

# Causal inference in the relative survival framework

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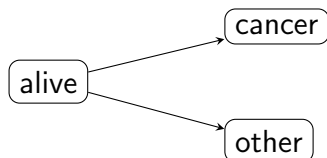
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# Background

In cancer epidemiology, competing risks are common.



We can estimate one of the following:

- Cause-specific mortality
- Excess mortality

Problems may arise by the inaccuracy or non-availability of the cause of death information.

We focus on estimating excess mortality that does not require the information on cause of death.

# Excess mortality and Relative survival

## Excess mortality

excess mortality = all-cause mortality - expected mortality

$$\lambda(t) = h(t) - h^*(t)$$

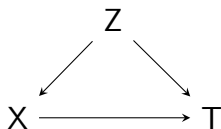
## Relative survival

relative survival ratio =  $\frac{\text{all-cause survival proportion}}{\text{expected survival proportion}}$

$$R(t) = \frac{S(t)}{S^*(t)} \quad S(t) = S^*(t)R(t)$$

The expected survival is considered to be known and is obtained from available life tables on a comparable population.

# Average causal effects



- Let  $T_i^0$  and  $T_i^1$  be the outcomes (time-to-event) that would be observed if patient  $i$  was unexposed ( $X = 0$ ) or exposed ( $X = 1$ ) respectively (potential outcomes).
- For each patient, one of the potential outcomes is counterfactual and will never be realised.
- Estimation therefore focuses on estimating averages, such as the average causal difference

$$E(T^1) - E(T^0)$$

$$Pr(T^1 < t) - Pr(T^0 < t)$$

# Standardised survival function

The standardised survival function at  $X = x$ , is defined as

$$E[S(t|X = x, Z)] = \frac{1}{N} \sum_{i=1}^N \hat{S}(t|X = x, z_i).$$

- Fit a survival model, such as Cox model or flexible parametric model.
- Obtain survival predictions for each individual in the population.
- Calculate an average of the survival predictions.

## Difference of standardised survival functions

If interested in **all-cause** survival:

$$E [S(t|X = 1, Z)] - E [S(t|X = 0, Z)] \\ E [S^*(t|X = 1, Z)R(t|X = 1, Z)] - E [S^*(t|X = 0, Z)R(t|X = 0, Z)]$$

If interested in **cancer-related** survival:

$$E [R(t|X = 1, Z)] - E [R(t|X = 0, Z)]$$

- This difference will refer to a hypothetical world where the cancer of interest is the only possible cause of death.
- However, it can be a very useful measure when comparing populations!
- Average causal difference under assumptions: conditional exchangeability, consistency, positivity, no interference, well-defined exposures.

## Standardising to a subset of the study population

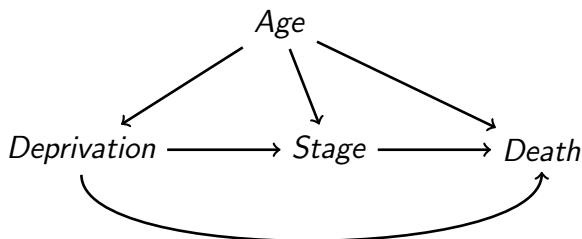
By standardising to the covariate distribution of the exposed:

$$E [R(t|X = 1, Z_1)] = \frac{1}{N_1} \sum_{i=1}^{N_1} \hat{R}(t|X = 1, z_{1i})$$

$$E [R(t|X = 0, Z_1)] = \frac{1}{N_1} \sum_{i=1}^{N_1} \hat{R}(t|X = 0, z_{1i}),$$

where  $N_1$  is the proportion of patients within the exposed.

# Mediation analysis



*How much of the differences between deprivation groups can be explained by differences at the stage distribution?*

*How many deaths could be avoided if the most deprived patients had the same relative survival as the least deprived? or same stage distribution?*



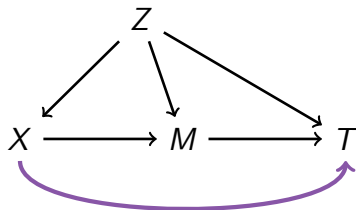
# Natural Direct Effects

## Natural Direct Effects

For each level of  $Z$ :

$$NDE = E [R(t|X = 1, M_1)] - E [R(t|X = 0, M_1)]$$

Note here the exposed have their own mediator distribution, but the unexposed have the mediator distribution of the exposed.



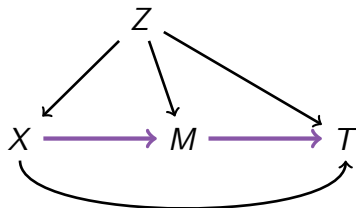
# Natural Indirect Effects

## Natural Indirect Effects

For each level of  $Z$ :

$$NIE = E [R(t|X = 1, M_1)] - E [R(t|X = 1, M_0)]$$

Gives the effect of the mediator in the exposed.



## Real world impact - Avoidable deaths

- The predicted number of deaths within strata  $Z = z$  for exposure  $X = 1$ :

$$D_1(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 1)R(t|X = 1, M_1)])$$

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$$D_1(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 1)R(t|X = 1, M_1)])$$

- The expected number of deaths if the exposed group had the same survival as the unexposed is

$$D_{1|0}(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 0)R(t|X = 0, M_1)])$$

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- The expected number of deaths if the exposed group had the same **(relative)** survival as the unexposed is

$$D_{1|0}(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 0)R(t|X = 0, M_1)])$$

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- The avoidable deaths are therefore,

$$AD(t|X = 1, M_1) = D_1(t|X = 1, M_1) - D_{1|0}^c(t|X = 1, M_1)$$

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$$D_1(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 1)R(t|X = 1, M_1)])$$

- The expected number of deaths if the exposed group had the same **(relative)** survival as the unexposed is

$$D_{1|0}(t|X = 1, M_1) = N_{Z=z, X=1} \times (1 - E[S^*(t|X = 0)R(t|X = 0, M_1)])$$

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- The avoidable deaths are therefore,

$$AD(t|X = 1, M_1) = D_1(t|X = 1, M_1) - D_{1|0}^c(t|X = 1, M_1)$$

- To get the total avoidable deaths we sum over all strata,  $Z$ .

$$AD(t, X = 1, M_1) = \sum_{Z_1} AD(t|X = 1, M_1)$$

- We analysed data on colon cancer in England.
- Fitted a flexible parametric survival model that uses restricted cubic splines to model the baseline excess hazard:
  - Age at diagnosis as a continuous and non-linear variable
  - Time-dependent effects for deprivation and age
  - Interaction between deprivation and age
- Performed a period analysis with a period window from 2012 to 2013 to ensure that the estimates are accurate for those recently diagnosed.
- Multiple imputations methods for missing data at stage at diagnosis.
- Background mortality was incorporated in the model using life tables.

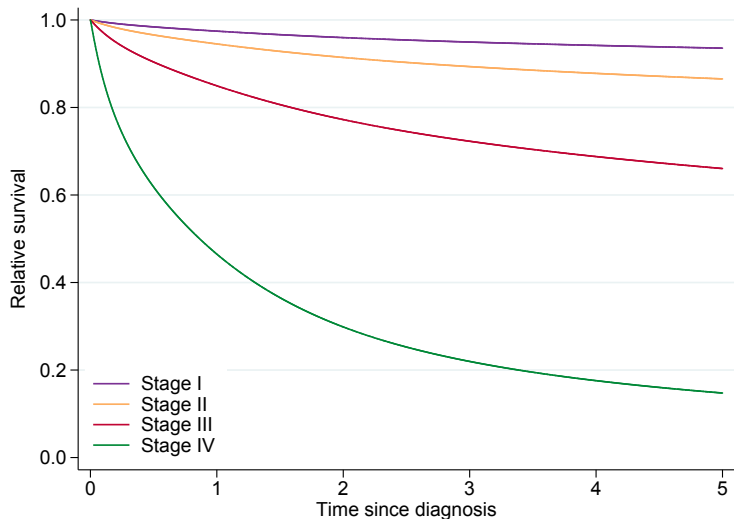


## Relative survival and stage by deprivation (females)

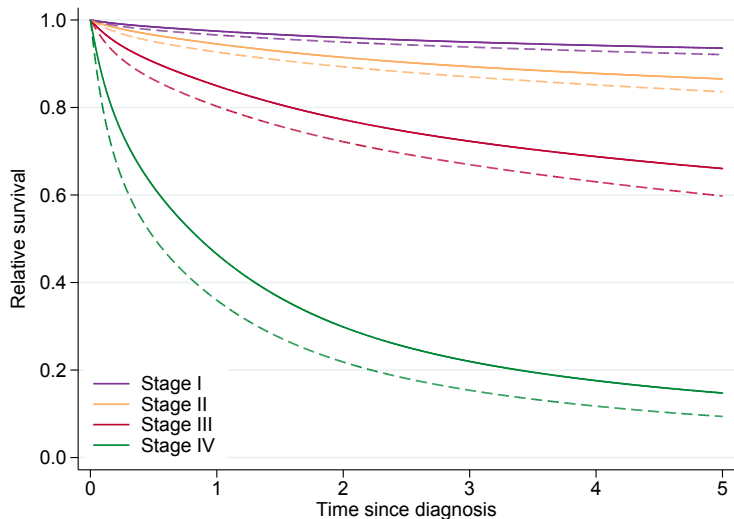
Age at Diagnosis	5-year Relative Survival	
	<i>Least deprived</i>	<i>Most deprived</i>
55	68.2	59.6
<b>65</b>	<b>66.5</b>	<b>56.6</b>
75	62.8	54.3
85	46.0	41.0

Stage at Diagnosis	Proportion of patients	
	<i>Least deprived</i>	<i>Most deprived</i>
I	14.0	13.4
II	29.9	30.5
III	28.2	25.9
<b>IV</b>	<b>27.9</b>	<b>30.2</b>

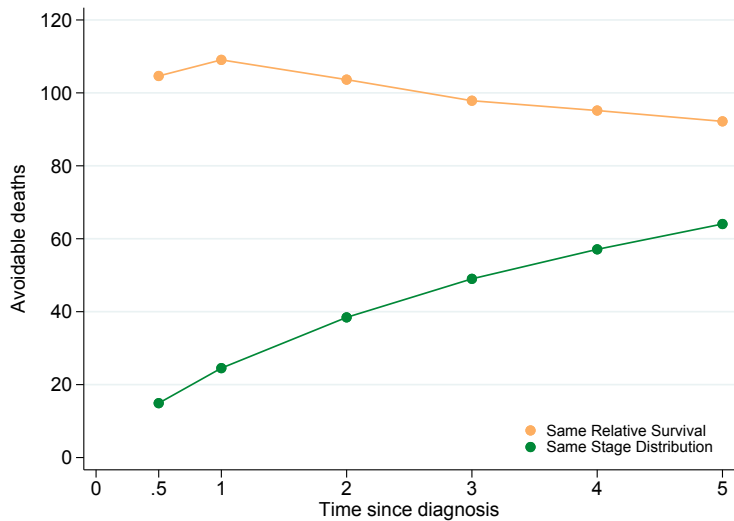
## Relative survival by stage - least deprived



# Relative survival by stage - least and most deprived



# Avoidable deaths



# Conclusions

- Relative survival controls for different background mortality across populations and enables fair comparisons.
- Avoidable deaths help quantify the impact of eliminating cancer inequalities in the real world.
- By applying a causal approach in cancer epidemiology we can make conclusions on the causal structure of variables and understand the underlying determinants of inequalities.
- Additional assumptions need to hold for the mediation analysis framework.
- Future work:
  - Model the mediator instead of stratification.
  - Extend software to estimate avoidable deaths at more than one time point.

# Selected References



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